

Bayesian Machine Learning

May 2022 - François HU https://curiousml.github.io/











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Lecture 1 : Bayesian inference

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$$P(y|x, X^{(train)}, Y^{(train)}) \approx \frac{1}{n} \sum_{i=1}^{n} P(y|x, \theta_i)$$

if we can sample $\theta_1, \ldots, \theta_n \sim P(\theta | X^{(train)}, Y^{(train)})$

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Lecture 2 (and 3) : M-step of EM-algorithm

$$\max_{\theta} \mathbb{E}_T \left[\log P(X, T | \theta) \right]$$

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if we can sample $\theta_1, \ldots, \theta_n \sim P(\theta | X^{(train)}, Y^{(train)})$

if we can sample $T_1, \ldots, T_n \sim T$

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with $P(\theta | X^{(train)}, Y^{(train)}) = \frac{p(Y^{(train)} | X^{(train)}, \theta) \cdot P(\theta)}{P(\theta)}$ easy : model output + prior fixed by us difficult : as always ... constant $\text{if we can sample } \theta_1, \dots, \theta_n \sim P(\theta \,|\, X^{(train)}, \dots, Y^{(train)}) \\ \text{for } N^{(train)} \\ \text{for } N^{$ if we can sample with $.., T_n \sim T$

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Starting point : we know how to simulate a pseudo-random uniform $U \sim \mathcal{U}(0,1)$

For « usual » distributions : both discrete and continuous r.v. can be sampled thanks to the uniform distribution

In practice (with python) we can easily sample them (via scipy and numpy for example)

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Otherwise : if there isn't an analytical way to sample it then **Rejection sampling** algorithm.

Assumption : we can compute distribution's pdf P and sample from an auxiliary distribution Q s.t. $P \le \text{const} \times Q$

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- **1.** generate sample $x_i \sim Q$ (auxiliary distribution)
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- **3.** if $u \leq P(x_i)$ then accept x_i else reject.

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if the \triangleleft gaps \rightarrow between P and Q are too large, we reject most of the sample



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Monte Carlo sampling : generates independent samples from the probability distribution in order to estimate an expected value



where $X_1, \ldots, X_n \sim P$ i.i.d

2. Markov Chain Monte Carlo **Definition : Markov Chain**

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Markov Chain : generates a sequence of r.v. where the *next* variable is probabilistically dependent upon the current variable.

P is called **stationary** if $P(x') = \sum_{x \in \text{supp}(\mathbf{V})} T(x, x') \cdot P(x)$ $x \in supp(X)$

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Objective : Build a Markov Chain that converges to the target distribution P no matter the starting point

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Markov Chain Monte Carlo : Algorithms

Reminder : we want to sample $x^{(1)}, ..., x^{(n)} \sim P(x_1, x_2, ..., x_d)$

Remark : we denote $x^{(i)} := (x_1^{(i)}, \dots, x_d^{(i)})$; $x_{-j} = (x_1, \dots, x_{j-1})$

Gibbs Sampling Algorithm

- Hypothesis : The conditional $P(x_i | x_{-i})$ can be sampled
- Initialisation : $x^{(0)} = (0, ..., 0)$ or random values
- **Repeat**: _

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$$x_{j+1}, x_{j+1}, \dots, x_d$$
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sometimes it can converge slowly to the desired distribution sometimes Gibbs samples can be too correlated

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- Initialisation : $x^{(0)} = (0, ..., 0)$ or random values
- **Repeat** :

sample $x^{(i)} = (x_1^{(i)}, \dots, x_d^{(i)})$ based on $x^{(i-1)} = (x_1^{(i-1)}, \dots, x_d^{(i-1)})$

for each position, $x_k^{(i)} \sim P(x_1 \mid x_{1:k-1}^{(i)}, x_{k+1:d}^{(i-1)})$

sometimes it can converge detropolis-Hasti Metropolis-Hasti Use a variant Gibbs sampling : Metropolis-Hasti Dise a variant Gibbs sampling in the too correlated

)
$$x_{1}, x_{j+1}, \dots, x_{d}$$
; $x_{m:n} = (x_m, x_{m+1}, \dots, x_n)$



Reminder : we want to sample $x^{(1)}, ..., x^{(n)} \sim P(x_1, x_2, ..., x_d)$

Remark : we denote $x^{(i)} := (x_1^{(i)}, \dots, x_d^{(i)})$; $x_{-j} = (x_1, \dots, x_{j-1})$

Metropolis-Hastings Algorithm

- Hypothesis : Let $P = \hat{P}/\text{const}$ where \hat{P} can be calculated and let Q be an auxiliary distribution we can sample from
- Initialisation : $x^{(0)} = (0, ..., 0)$ or random values
- **Repeat** :

)
$$x_{j+1}, x_{j+1}, \dots, x_d$$
; $x_{m:n} = (x_m, x_{m+1}, \dots, x_n)$

Reminder : we want to sample $x^{(1)}, ..., x^{(n)} \sim P(x_1, x_2, ..., x_d)$

Remark : we denote $x^{(i)} := (x_1^{(i)}, \dots, x_d^{(i)})$; $x_{-j} = (x_1, \dots, x_{j-1})$

Metropolis-Hastings Algorithm

- Hypothesis : Let $P = \hat{P}/\text{const}$ where \hat{P} can be calculated and let Q be an auxiliary distribution we can sample from
- Initialisation : $x^{(0)} = (0, ..., 0)$ or random values
- **Repeat** : _

sample a candidate $x^{(i)} \sim Q(x^{(i)} | x^{(i-1)}) = (example of auxiliary distribution) <math>\mathcal{N}(x^{(i-1)}, \sigma^2 I)$

)
$$x_1, x_{j+1}, \dots, x_d$$
; $x_{m:n} = (x_m, x_{m+1}, \dots, x_n)$

Reminder : we want to sample $x^{(1)}, ..., x^{(n)} \sim P(x_1, x_2, ..., x_d)$

Remark : we denote $x^{(i)} := (x_1^{(i)}, \dots, x_d^{(i)})$; $x_{-j} = (x_1, \dots, x_{j-1})$

Metropolis-Hastings Algorithm

- Hypothesis : Let $P = \hat{P}/\text{const}$ where \hat{P} can be calculated and let Q be an auxiliary distribution we can sample from
- Initialisation : $x^{(0)} = (0, ..., 0)$ or random values
- **Repeat** : _

sample a candidate $x^{(i)} \sim Q(x^{(i)} | x^{(i-1)}) = (example)$ with acceptance probability : min $\left(1, \frac{Q(x^{(i-1)} | x^{(i)})}{Q(x^{(i)} | x^{(i-1)})}\right)$

)
$$x_1, x_{j+1}, \dots, x_d$$
; $x_{m:n} = (x_m, x_{m+1}, \dots, x_n)$

e of auxiliary distribution)
$$\mathcal{N}(x^{(i-1)}, \sigma^2 I)$$

 $\stackrel{(i)}{\to} \times \hat{P}(x^{(i)})$ accept $x^{(i)}$ as an sample from P

Reminder : we want to sample $x^{(1)}, ..., x^{(n)} \sim P(x_1, x_2, ..., x_d)$

Remark : we denote $x^{(i)} := (x_1^{(i)}, \dots, x_d^{(i)})$; $x_{-j} = (x_1, \dots, x_{j-1})$

Metropolis-Hastings Algorithm

- Hypothesis : Let $P = \hat{P}/\text{const}$ where \hat{P} can be calculated and let Q be an auxiliary distribution we can sample from
- Initialisation : $x^{(0)} = (0, ..., 0)$ or random values
- **Repeat** : _

sample a candidate $x^{(i)} \sim Q(x^{(i)} | x^{(i-1)}) = (example)$ with acceptance probability : min $\left(1, \frac{Q(x^{(i-1)}|x^{(i)})}{Q(x^{(i)}|x^{(i-1)})}\right)$

)
$$x_1, x_{j+1}, \dots, x_d$$
; $x_{m:n} = (x_m, x_{m+1}, \dots, x_n)$

e of auxiliary distribution)
$$\mathcal{N}(x^{(i-1)}, \sigma^2 I)$$

 $\hat{P}(x^{(i)})$ accept $x^{(i)}$ as an sample from P

Reminder : we want to sample $x^{(1)}, ..., x^{(n)} \sim P(x_1, x_2, ..., x_d)$

Remark : we denote $x^{(i)} := (x_1^{(i)}, ..., x_d^{(i)})$; $x_{-j} = (x_1, ..., x_{j-1})$

Metropolis-Hastings Algorithm

- Hypothesis : Let $P = \hat{P}/\text{const}$ where \hat{P} can be calculated and let Q be an auxiliary distribution we can sample from
- Initialisation : $x^{(0)} = (0, \dots, 0)$ or random values
- **Repeat** : _

rho = 0.9, tau = 0.001 rho = 0.9, tau = 1 rho = 0.9, tau = 5 0.50 0.25 0.00



)
$$x_1, x_{j+1}, \dots, x_d$$
; $x_{m:n} = (x_m, x_{m+1}, \dots, x_n)$



3.b. MCMC vs VI pros and cons

MCMC

Pros:

- Useful when the posterior is intractable
- Asymptotically exact
- Suited to small / medium dataset

Cons:

- Usually slower than alternatives (VI)
- Can generates dependant samples from the distribution

VI (see lecture 3)

Pros : - Useful when the posterior is intractable - Suited to large dataset

Cons : - Can never generate exact result







- dataset

- distribution